DOCUMENT RESUME

ED 275 698 TM 860 575

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TITLE Assessing the Dimensionality of the GMAT Verbal and

Quantitative Measures Using Full Information Factor

Analysis.

INSTITUTION Educational Testing Service, Princeton, N.J.

SPONS AGENCY Graduate Management Admission Council, Princeton,

NJ.

REPORT NO ETS-RR-86-13

PUB DATE Mar 86 NOTE 21p.

PUB TYPE Reports - Research/Technical (143)

EDRS PRICE MF01/PC01 Plus Postage.

DESCRIPTORS Business Administration Education; College Entrance

Examinations; *Factor Analysis; *Factor Structure; Graduate Study; Higher Education; *Item Analysis;

*Latent Trait Theory; Mathematical Models;

Mathematics Tests; *Maximum Likelihood Statistics:

Verbal Tests

IDENTIFIERS *Graduate Management Admission Test; *TESTFACT;

Unidimensionality (Tests)

ABSTRACT

When the three-parameter logistic model and item response theory are used to analyze Graduate Management Admission Test (GMAT) data, there are problems with the assumption of unidimensionality. Linear factor analytic models, exploratory factor analysis programs, and the comparison of item parameter estimates for heterogeneous and homogeneous subsets also present difficulties. A new method of assessing the dimensionality of binary data is now available. TESTFACT is a computer program which can be used to perform full information factor analysis, using the marginal maximum likelihood method to estimate reparameterized discrimination and difficulty parameters for multidimensional item response models. The lower asymptote for each item is treated as a known constant whose value is input by the program user. TESTFACT allows a stepwise factor analysis to be performed. First a one-factor solution is performed, then a two-factor solution. The difference between chi squares for the two solutions is used to test whether the added factor is statistically significant. When TESTFACT was applied to both quantitative and verbal GMAT items, a dominant first factor and two additional, considerably weaker, factors were found. (Author/GDC)



ASSESSING THE DIMENSIONALITY OF THE GMAT VERBAL AND QUANTITATIVE MEASURES USING FULL INFORMATION FACTOR ANALYSIS

Neal Kingston

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Educational Testing Service Princeton, New Jersey March 1986



Assessing the Dimensionality of the GMAT Verbal and Quantitative Measures Using Full Information Factor Analysis

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This Research was Sponsored by the Graduate Management Admission Council

March 1986



The consultation and review of Lawrence Hecht, Robert Mislevy, Nancy Petersen, and Martha Stocking is gratefully acknowledged. Also, thanks to Peter Pashley for getting TESTFACT up and running. The opinions expressed herein are those of the author and do not neccesarily reflect those of Educational Testing Service, the Graduate Management Admission Council, nor any of the reviewers and consultants.

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INTRODUCTION

Kingston, Leary, and Wightman (1985) explored the applicability of item response theory methods for the Graduate Management Admission Test. In their report they assessed the fit of the three-parameter logistic model to GMAT data in a number of ways. In particular, they assessed the appropriateness of the unidimensionality assumption using two methods: 1) analysis of previous exploratory factor analyses (i.e., Swinton & Powers, 1981) and 2) comparison of item parameters estimated both for homogeneous and heterogeneous subsets of items. These analyses indicated that the Verbal and Quantitative measures of the GMAT each probably has two major dimensions, and possibly a number of minor dimensions.

Each of these types of analyses, however, has a theoretical drawback for assessing dimensionality in an IRT framework. Although the three-parameter logistic model assumes unidimensionality, it does not require that the dimension be linearly related to the scored item responses (right vs. wrong) from which the trait is drawn.

Thus, linear factor analytic models might present a misleading picture. Also, commonly available exploratory factor analysis programs do not indicate whether the factors are statistically significant. The comparison of item parameter estimates for items calibrated both in homogeneous and heterogeneous subsets also provided no indication as to whether or not the differences were statistically significant.



A new method of assessing the dimensionality of binary data is now available. TESTFACT (Wilson, Wood, & Gibbons, 1984) is a computer program which can be used to perform full information factor analysis (Bock, Gibbons, & Muraki, 1985). Full information factor analysis, as implemented in TESTFACT, uses the marginal maximum likelihood method (Bock & Aitken, 1981) to estimate (reparameterized) discrimination and difficulty parameters for multidimensional IRT models. The lower asymptote for each item is treated as a known constant whose value is input by the program user. TESTFACT allows a stepwise factor analysis to be performed. First a one-factor solution is performed, then a two-factor solution. The difference between chi squares for the two solutions is used to test whether the added factor is statistically significant. A third, fourth, or more factors can be added, but computation time and expense increases exponentially with the number of factors.



RESEARCH DESIGN

Description of the Test

The data presented in this report were obtained from the Graduate Management Admission Test, which reports Verbal, Quantitative, and Total scaled scores. The test consists of eight separately timed sections, two of which were not used as part of this study. The reported scores are derived from six of the sections. The Verbal measure consists of 85 items administered in three sections: reading comprehension (25 items), sentence correction (25 items), and analysis of situations (35 items). The Quantitative measure includes 65 items: two sections of problem solving items (40 items total) and one section of data sufficency items (25 items). Examples of all item types are available in the GMAT Bulletin of Information (GMAC, 1985).

Data Collection

Data analysis was based on random sample of 5,000 examinees who took GMAT form 3FBS1 at the January 1983 administration.

<u>Analysis</u>

A scored item tape was created for input into the analyses.

Although the GMAT is scored operationally as number right minus one-quarter number wrong, for the purpose of this study all items were scored 0 if wrong and 1 if right. For the calculation of tetrachoric correlations there are three options: 1) delete all examinees who omitted any item, 2) do pairwise calculation of tetrachorics, deleting examinees who omitted one of the pair of items, and 3) code omitted items as wrong. In addition, TESTFACT allows all omitted items



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following the last item to which the examinee responded to be treated as not presented. This was not done as such treatment would make difficult the finding of any speed factor, should one exist.

A one factor, two factor, and three factor full information factor analysis was run separately for the GMAT Verbal and Quantitative measures. For the verbal analysis, the one- and two-factor solutions converged readily, but for the three-factor solution a Bayesian prior distribution was set on the item parameters in order to aid convergence. A prior was used for all three quantitative analyses.





RESULTS

<u>Verbal</u>

Table 1 presents the results for the full information factor analysis of the verbal items. Both the second and third factors were clearly statistically significant at any commonly accepted level (beyond the .0001 level). It appears likely that additional factors might also have been statistically significant if a higher factor model had been run. It should be noted, however, that the percent of score variance explained for the three factors was 21.3, 3.9, and 3.2, respectively, indicating that the overwhelming plurality of the explained variance was determined by a dominant first factor.

Insert Table 1 About Here

A promax rotation of the three factor solution (Hendrickson & White, 1964) indicated that the factors were fairly highly correlated $(r_{12} = .59; r_{13} = .65; and r_{23} = .59)$. The two factor solution with promax rotation yielded a .58 correlation between the factors.

Table 2 presents the factor loadings for the two and three factor solutions. If items were assigned to factors based on the factor on which they had the highest loading, then for the two factor solution, all reading comprehension and sentence correction items would have been assigned to the first factor. All but two of the analysis of situations items would have been assigned to the second factor. These two had slightly higher loadings on the first factor.



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Insert Table 2 About Here

•••••

One way to test the interpretability of this solution is to compare the correlation between the promax rotated factors with the correlation between formula scores on the analysis of situations section and the combined sentence correction and reading comprehension sections. Since the factor scores contain some information from the items constituting the other factor and since the relative contribution of items that load heavily on the appropriate factor vary with the strength of their relationship to the factor, factor scores tend to be more reliable than formula-scores. This tends to increase the correlations between factors so that they are higher than the correlations between the observed scores from which those factors were derived. To the extent that the correlation between a pair of observed scores is close to the correlation between the corresponding pair of factor scores, then the item types constituting the observed scores define the factors. Since the factors correlate .58 and the combined sentence correction and reading comprehension section scores correlate .56 with analysis of situations scores, it is clear that it is item type that defines these two factors.

Assigning items to factors for the three-factor solution yields the same clearly defined first factor. All reading comprehension and sentence correction items loaded most heavily on the first factor. One analysis of situations item also loaded most heavily on the first



factor, although its loadings on the second and third factor were similar in magnitude. The other analysis of situations items all loaded most heavily on either factor two (15 items) or factor three (19 items).

Five items had loadings of .7 or greater on factor two. Each of these items had the correct answer in the "C" position. Nine of the remaining ten items that loaded most heavily on factor two were keyed either "C" (4 items) or "E" (5 items). The remaining item had a key of "B", but its loading was low on all three factors, indicating that the care variance that was contributed was largely specific to that item and not to any common factors.

Six of the items had loadings of .7 or greater on factor three. Five of these items had an answer key of "A", and the last had a key of "D". All of the remaining 13 items that loaded most heavily on factor three were keyed "B" (6 items) or "D" (6 items).

Analysis of situations item share a common set of response options. "A" always indicates that the given item is a major objective.
"B" indicates a major factor, "C" a minor factor, "D" a major assumption, and "E" an unimportant issue. Thus, it appears clear that the second and third factors are analysis of situations answer key-factors. Swinton and Powers (1981) found similar factors using classical exploratory factor analysis. Previous research on a similar item type called "analysis of explanations," which had previously been used in the Graduate Record Examinations Aptitude Test, indicates that such answer key factors are not uncommon (Kingston & Dorans, 1985; Swinton & Powers, 1980).



Quantitative

Table 3 presents the results of the full information factor analysis of the quantitative items. The first three factors are statistically significant at beyond the .0001 level. It is possible that additional significant factors might exist. The first factor is clearly dominant, explaining 33.7 percent of the observed score variance, compared to 3.5 percent for the second factor and 1.3 percent for the third factor.

Insert Table 3 About Here

Two-factor solution. A promox rotation of the two-factor solution showed that the factors correlated .72. Table 4 presents the factor loadings for the two- and three-factor solutions. For the two-factor solution, 37 of the 65 items loaded more heavily on the first factor and the remaining 28 items loaded more heavily on the second factor. Many items loaded about the same with each of the two factors. Some problem solving items loaded more heavily on the first factor and other loaded more heavily on the second. Similarly, different data sufficiency items loaded on each of the factors.

Insert Table 5 About Here

GMAT quantitative items can be fit into a three-way classification scheme: item type (problem solving or data sufficiency) mathematics type (arithmetic, algebra, or geometry) and problem type (pure or



applied). Examination of the content of the quantitative items shed some light on the constructs or abilities underlying the factors. Table 5 presents the relationship between item content and factor loadings.

For the two-factor solution, all pure geometry items and most pure algebra items loaded more heavily on the first factor. Most applied algebra items, and applied and pure arithmetic items loaded more heavily on the second factor. The seven applied geometry items were split across the two-factors. Thus, the first factor might best be called a pure geometry and algebra factor, and the second factor might be called an arithmetic and applied algebra factor.

While the interpretation of the quantitative factors is not quite as clear as the interpretation of the verbal factors, this is not surprising. The first quantitative factor is relatively larger than the second quantitative factor compared to the corresponding relationship between verbal factors, and the correlation between the two quantitative factors is very high, about .72. In short, the two factor solution suggests that the bulk of explainable common variance in Quantitative scores for the GMAT population is due primarily to differences in the ability to perform geometric and algebraic manipulations in non-applied settings. Most of the remaining common variance cuts across quantitative item content, with the rest explained by variance due to the idiosyncratic item content or measurement error.



Three-factor solution. A promax rotation of the three factor solution showed that the factors were highly correlated $(r_{12} - .61; r_{13} - .68; and r_{23} - .73)$. As with the two factor solution, there did not seem to be any consistent relationship between item type and factor loadings. Item content did appear to explain the factors that were found.

All pure geometry items and most pure algebra items had their largest loading the first-factor. Fifteen of the 16 applied arithmetic, problem solving items loaded most heavily on factor two but six of the eight applied arithmetic data sufficiency items had their largest loading on factor three. The other items were scattered among the three factors. Thus, the first factor can be characterized as pure geometry and pure algebra, the second factor as applied arithmetic, problem solving, and the third factor as applied arithmetic, data sufficiency. The latter two factors, however, were very weak, with the third factor explaining only about one percent of the score variance. Also, all three factors were fairly highly correlated.

In summary, both the two- and three-factor solutions supported the existance of a single dominant quantitative factor.



DISCUSSION AND CONCLUSION

For both the verbal items and quantitative items a dominant first factor and two additional, considerably weaker factors were found. This confirms the findings of Kingston, Leary, and Wightman (1985), but is a stronger analysis in that it is a direct nonlinear factor analytic approach as compared to the earlier approach based primarly on the comparision of item parameters estimated from heterogeneous and homogeneous subsets of items. Also, the use of full information factor analysis provided a statistical test of the factor model that confirmed the presence of multiple factors.



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Table 1
Full Information Factor Analysis of GMAT Verbal Items

Number Factors in Solution	Latent Root*	Percent Variance Explained*	2 X	df_	2 X Change**	df <u>Change</u> **	P(χ ²) <u>Change</u> ***
1	24.0	21.3	406,006	4,829	• • •		
2	3.7	3.9	402,533	4,745	3,473	84	.0000
3	2.6	3.2	399,794	4,662	2,734	83	.0000

previous model x^2 where x^2 was a substitute of change in x^2 under null hypothesis



^{*}Latent roots and variance explained are from three factor

^{**}Difference between χ^2 or degrees of freedom for this model and previous model

	Teble 2 Verbel itas Laadings						
	Ina-Pacta 	r saluting	<u> 1</u>	erea-factur as	lution.		
RC 1	0.374	-0.047	0.4		-0.070		
RČ Ž RC 3	0. 61 2 0. 255	-0.012 0.007	0.6 0.2		-0.050 -0.005		
NG 4	0.474	-0.014	0.7	27 0.043	-0.111		
AC S	0.200	0.021	0.2		-0.030		
RC & RC 7	0.490 0.530	0.016 -0.045	0.5 0.5		-0. 0 5 7 -0. 007		
RC 0	0.512	-0.074	0.5	31 -0.043	-0.074		
RC 9 RC 10	0.334 0.436	-0.023 -0.213	0.3 0.4		-0.040 -0.111		
RC 11	0.430	-0.049	0.4	-0.000	-0.044		
RC 12 RC 13	0.447 0.402	0.030 -0.046	0.4 0.4		-0.017 -0.000		
RC 14	0.470	-0.023	9.4	98 0.003	-0.035		
RC 10	0.256	-0.004	0.2		0.030		
RC 16 RC 17	0.143 0.663	-0.004 -0.092	0.1 0.6		0.033 -0.019		
RC 10	0.745	-0.129	0.7	-0.116	0.007		
RC 19 RC 20	0.664 0.657	-0.100 -0.047	0.		0.00 5 -0.013		
RC 21	0.444	-0.091	0.4	-0.070	0.007		
AC 22 AC 23	. 0.465	-0.121 -0.043	0.4 0.5	9 -0.101	-0.020 0.011		
NC 24	0.414	-0.013	0.5	74 -0.099	0.100		
RC 25	0.733	-0.152	0.4		0.001		
A0 1 A0 2	0.070 0.267	0. 235 0. 307	0.0 0.2		0.355 0.342		
Ã0 3	0.230	0.371	0.2	64 0.314	0.003		
A0 4 A0 0	0.027 0.112	0.349 0.439	-0.0 •.0		0.363 0.299		
A0 6	0.031	0.371	-0.0	25 0.092	0.407		
A0 7 A0 0	0.249 -0.211	0.374 0.737	0.0 -0 .0		0.025 -0.041		
ÃO Ý	0.331	0.479	0.0	4 -0.047	0.045		
A0 10	-0.201	0.706	-0.0		-0.011		
A0 11 A0 12	0.1 99 0.356	0. 399 0. 4 0 3	0.2 0.2		0.100 0.534		
A0 13 A0 14	-0.0 06 -0.231	0. 902 0. 724	0.0	07 0.540	-0.077 -0.204		
A0 15	0.078	0. 451	0.6		0.310		
A0 16	-0.273	0.052	-0.1	12 0.926	-0.133		
A0 17 A0 10	0.04 0 0.077	0.097 0.415	0.1 -0.0		-0.097 0.769		
A0 17 A0 20	0.220 0.126	0.309 0.097	0.2 0.1	00 0.221	0.250 0.005		
A0 21	0.149						
A0 22	0.100	0.2 0 5 0.320	0.0 0.1	03 0.173	0.533 0.214		
A0 23 A0 24	-0.032 0.111	0.500 8.303	0.0 0.1		0.00 7 0.002		
A0 28	0.064	0.140	0.0		0.030		
A0 26	-0.304	0.923	-0.1		-0.172		
A0 27 A0 20	0.052 0.072	0.321 0.400	0.0 -0.0		0.07 6 0.017		
A0 27 A0 30	0.234 0.017	0.432 0.010	0.1 0.0	91 0.178	0.362 0.151		
A0 31							
A0 32	0.077 0.371	0.467 0.346	-0.1 0.2	15 -0.057	0.002 0.435		
AS 33 AO 34	0.166 -0.107	0.332 0.425	1.0 -1.1		0.4 5 3 0.074		
ÃÔ 35	0.199	0.204	0.0		0.045		
8C 1 8C 2	0.3 05 0.24 0	-0.047	1.3		-0.021		
ec 3	0.381	-0.004 -0.089	0.2 0.3	72 0.005	-0.014 -0.004		
9C 4 9C 5	0.307 0.3 6 5	0.074 0.010	0.3 0.3		0.022 -0.032		
•C A	0.410	-0.047	0.3		-0.027		
8E 7 9C •	0.440	-0.000	0.4	0.026	-0.039		
90 9	0.280 9.474	0.044 -0.046	0.2 0.4		0.020 =0.032		
SC 10	0.440	0.023	0.4		0.003		
8C 11 9C 12	0.434 0.376	0.078 -0.007	0.4		0.035		
8C 13	0.833	-0.075	0.3 0.5	31 -0.076	0.010 0.002		
SC 14 9C 18	0.293 0.262	0.107 0.100	0.2 0.2		0.070 -0.005		
9C 16	0.363	0.003	0.3		-0.017		
● C 17	0.434	0.010	0.4	30 0.017	0.001		
86 10 86 19	0.42 8 0.522	0.0 53 0.017	0. 4 0. 5		0.024 0.070		
SĆ 20	9.442	-0.016	1.1		0.111		
SC 21	0.521	-0.044	10 !!		-0.010		
9C 22 9C 23	0.364 0.688	-0.043 -0.050	18	14 0.007	0.010 -0.046		
BC 24 OG 28 Die bill erkel in einerheide bill bisekkling bill broken.	0.457	0.033	0.4	24 -4.493	8.184		

Table 3
Full Information Factor Analysis of GMAT Quantitative Items

Number Factors in Solution	Latent Root*	Percent Variance Explained*	2 X	<u>df</u>	2 X <u>Change</u> **	df <u>Change</u> **	P(χ^2) Change
1	26.9	33.7	245,107	4,868	•-•		
2	2.0	3.5	244,006	4,804	1,101	64	.0000
3	1.4	1.3	243,339	4,741	668	63	.0000



^{*}Latent roots and variance explained are from three factor solution

^{**}Difference between χ^2 or degrees of freedom for this model and previous model

^{***}Probability of change in χ^2 under null hypothesis

Table 4 Quantitative Itee Loadings Two-Factor Bolution

		140-	. acre. adiation					
	Content		Two-Fector eclution		Three-Fector Bolution			
itee*	Clessification	_1_	2		_2_			
P1 1	Alashas - Aven							
P1 1 P1 2	Algebre - Pure Aritheetic - Pure	0.249 0.145	0.207	0.244	-0.104	0.434		
Pi 3	Becestry - Pure	0.149	0.3 8 1 0.231	0.143 0.278	0.200 0.011	0.149		
P1 4	Algebre - Applied	0.004	0.445	0.008	0.245	0.234 0.241		
P1 5	Aritheetic - Applied	0.144	V. 327	0.137	0.220	0.184		
	•••			*****	******	*****		
P1 6	Arithmetic - Applied	0.049	0.457	0.019	0.331	0.228		
P1 7	Aritheetic - Applied	-0.273	0.747	-0.278	0.499	0.341		
P1 9	Algebre - Applied Arithmetic - Applied	-0.045	0.733	-0.129	0.435	0.450		
P1 10	Aritheetic - Applied	-0.051 -0.044	0.409 0.537	-0.045	0.442	0.023		
	Misconnere - Whiting	-0.048	V. 337	-0.044	0.285	0.304		
P1 11	Aritheetic - Applied	0.115	0.541	0.087	0.479	0.181		
P1 12	Aritheetic - Pure	0.274	0.230	0.240	0.000	0.152		
P1 13	Aritheetic - Applied	-0.104	0.439	-0.115	0.411	0.090		
P1 14	Algebre - Pure	0.474	0.222	0.438	0.145	0.141		
P1 15	Aritheetic - Pure	0.340	0.384	0.308	0.259	0.248		
P1 14	Secretary - Applied	A 494	A 840	A 454				
P1 17	Geometry - Applied Algebra - Pura	0.471 0.311	0.207 0.104	0.431 0.289	0.289	0.014		
P1 18	Aritheetic - Applied	0.402	0.446	0.207	0.220 0.550	-0.043 0.044		
P1 19	Geometry - Pure	0.728	0.082	0.483	0.218	-0.044		
P1 20	Geometry - Applied	0.449	0.377	0.430	0.487	-0.055		
	• • •					******		
P2 1	Aritheetic - Applied	-0.147	0.449	-0.144	0.419	0.109		
P2 2 P2 3	Arithmetic - Applied	0.213	0.393	0.200	0.270	0.184		
P2 3 P2 4	Algebre - Pure Aritheetic - Pure	0.044	0.256	0.040	-0.007	0.297		
P2 5	Secentry - Pure	0.547 0.433	0.043 -0.004	0.551	-0.192	0.258		
	observe Para	V. 833	-0.008	0.430	-0.080	0.092		
P2 6	Secentry - Applied	0.292	0.341	0.304	0.238	0.134		
P2 7	Arithoetic - Applied	0.185	0.444	0.179	0.504	0.004		
P2 8	Aritheetic - Pure	-0.042	0.500	-0.100	0.221	0.378		
P2 •	Aritheetic - Pure	0.424	-0.024	0.424	-0.124	0.114		
P2 10	Algebre - Applied	0.005	0.504	0.029	0.453	0.083		
P2 11	Anishantia - Anniind		A 21A					
P2 12	Aritheetic - Applied Aritheetic - Applied	0.249 0.097	0.510	0.234	0.514	0.071		
P2 13	Aritheetic - Applied	0.379	0.488 0.378	0.047 0.349	0.449 0.411	0.141		
P2 14	Arithmetic/Algebre-Applied	0.585	0.310	0.531	0.344	0.028 0.044		
P2 15	Algebre - Pure	0.457	0.451	0.451	0.335	0.174		
	•			*****	******	******		
P2 16	Geometry - Applied	0.200	0.424	0.224	0.524	0.012		
P2 17 P2 18	Geometry - Pure	0.884	-0.019	0.891	0.181	-0.149		
P2 19	Algebre - Applied Algebre - Pure	0.432	0.120	0.424	0.214	-0.051		
P2 20	Aritheetic - Applied	0.949 0.396	-0.14 8 0.453	0.922	-0.055	-0.038		
	White	V. 376	V. 433	0.347	0.401	-0.071		
DS 1	Arithmetic - Applied	-0.380	0.740	-0.404	0.210	0.599		
DS 2	Algebre - Pure	0.134	0.272	0.130	0.085	0.217		
DB 3	Aritheetic - Pure	-0.034	0.584	-0.080	-0.023	0.487		
DS 4	Aritheetic - Applied	-0.047	0.515	-0.104	0.097	0.492		
D0 5	Geometry - Applied	0.235	0.351	0.193	0.081	0.354		
DS 4	Algebre - Applied	-0 190	A 500	4 554				
DS 7	Algebre - Applied	-0.172 0.134	0.529 0.397	-0.220 0.135	-0.024	0.441		
DS 8	Secentry - Applied	0.341	0.142	0.348	0.341 0.016	0.114 0.149		
DB 7	Aritheetic - Applied	-0.074	0.340	0.116	0.024	0.404		
DE 10	Aritheotic/Becastry-Applied	0.342	0.313	0.293	-0.004	0.431		
	• • • • • • • • • • • • • • • • • • • •							
DS 11 DS 12	Arithmetic - Applied	-0.054	0.370	-0.089	0.141	0.297		
DS 13	Secretry - Pure	0.478	0.234	0.424	0.053	0.203		
DB 14	Aritheetic - Applied Algebra - Applied	0.048 -0.131	0.300	-0.004	0.042	0.349		
DS 15	Secentry - Pure	0.279	0.524 0.175	-0.13 9 0.277	0.443	0.133		
		V16/7	4.170	4.477	0.174	0.054		
DS 14	Aritheetic - Applied	0.278	0.274	0.249	0.444	-0.120		
DS 17	Algebre - Pure	0.439	0.072	0.433	0.035	0.072		
DS 18	Algebre - Pure	0.847	-0.144	0.802	0.020	-0.073		
DS 19	Aritheetic - Pure	0.094	0.538	0.033	0.393	0.281		
DS 20	Algebre - Pure	0.274	0.195	0.140	-0.094	0.443		
DS 21	Geometry - Applied	0.502	0 004	A =44				
DS 22	Geometry - Pure	0.502	0.004 0.1 8 4	0.50 9 0.547	0.150 0.290	-0.147 -0.020		
DS 23	Algebre - Pure	r.435	0.310	0.347	0.273	-0.020 0.097		
DS 24	Arithoetic - Applied	0.435	0.349	0.349	0.407	0.077		
DS 25	Secentry - Pure	0.843	-0.022	0.783	0.200	-0.125		

ePi - Problem Solving, First section P2 - Problem solving, scond section D8 - Date Sufficiency



Table 5 Quantitative Measure Content Breakdown

		1	Wo-Factor Solution Factor with Larger Loading		Three-Factor Solution Factor with Largest Loading		
Content	Situation	Item-Type	1	2	1	2	<u>3</u>
Algebra	Pure	Problem Solving	4	1	4	•	1
Albegra	Pure	Data Dufficiency		ī	3	-	2
Algebra	Applied	Problem Solving	1	3	ì	2	ī
Algebra	Applied	Data Sufficiency		3	-	2	ī
Arithmetic	Pure	Problem Solving	2	4	3	1	2
Arithmetic	Pure	Data Sufficiency		2	-	ī	2
Arithmetic	Applied	Problem Solving	2	14	1	15	-
Arithmetic	Applied	Data Sufficiency		5	ō	2	6
Geometry	Pure	Problem Solving	4	-	4	_	_
Geometry	Pure	Data Sufficiency		-	4	-	-
Geometry	Applied	Problem Solving	2	2	2	2	_
Geometry	Applied	Data Sufficiency		ī	2	-	1

